


BIO4300 and ENV4300

# Stats fest 2007

## Analysis of variance

murray.logan@sci.monash.edu.au



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### Single factor ANOVA

- **Aims**
  - Description
    - Investigate differences between population means
  - Explanation
    - How much of the variation in response variable (Y) is explained by differences in the predictor variable (X)

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### Single factor ANOVA

- **Data**
  - Dependent (response) variable
    - Continuous
    - Normally distributed
  - Independent (predictor) variable
    - Categorical - factor
  - $n$  sampling units (replicates) per group (=level)

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### Single factor ANOVA

- Linear model
 
$$y = \beta_0 + \beta_1 X + \varepsilon$$

← Regression

$$y = \mu + \beta_1 G_1 + \beta_2 G_2 + \dots + \varepsilon$$

← ANOVA

ANOVA is equivalent to a multiple regression

a)		b)			
Y	A	Y	dummy <sub>1</sub>	dummy <sub>2</sub>	dummy <sub>3</sub>
2	G1	2	1	0	0
3	G1	3	1	0	0
4	G1	4	1	0	0
6	G2	6	0	1	0
7	G2	7	0	1	0
8	G2	8	0	1	0
10	G3	10	0	0	1
11	G3	11	0	0	1
12	G3	12	0	0	1

$y = \mu + \alpha_1 + \alpha_2 + \alpha_3 + \varepsilon$     Where  $\alpha_i$  is the effect of group 1, ...  
 $y = \mu + \alpha + \varepsilon$

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### Single factor ANOVA

- Linear model
 
$$y = \mu + \alpha_1 + \alpha_2 + \alpha_3 + \varepsilon$$
- Over parameterized
  - More parameters than there are groups
  - Cannot include all  $\alpha_i$ 's
  - Can only include  $p-1$  terms in model, plus the overall mean
  - Must redefine the  $\alpha$  terms

$y = \mu + \alpha + \varepsilon$

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### Single factor ANOVA

- Contrasts
  - redefine the  $\alpha$  terms
    - Treatment contrasts (default behaviour of R)
    - Can define own contrasts (Planned comparisons)

$y = \mu + \alpha_2 + \alpha_3 + \varepsilon$

over-parameterized design matrix				contrast matrix		model matrix			
Intercept	$\alpha_1$	$\alpha_2$	$\alpha_3$			Intercept	$\alpha_2^*$	$\alpha_3^*$	
$(\mu)$	$(G1)$	$(G2)$	$(G3)$	$G1$	$G2$				
$\begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix}$	$\begin{bmatrix} 1 \\ 1 \\ 0 \\ 0 \end{bmatrix}$	$\begin{bmatrix} 0 \\ 0 \\ 1 \\ 0 \end{bmatrix}$	$\begin{bmatrix} 0 \\ 0 \\ 0 \\ 1 \end{bmatrix}$	$\begin{bmatrix} \alpha_2^* \\ \alpha_3^* \end{bmatrix}$	$\begin{bmatrix} 0 \\ 0 \\ 1 \\ 1 \end{bmatrix}$	$\Rightarrow$	$\begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix}$	$\begin{bmatrix} 0 \\ 0 \\ 1 \\ 0 \end{bmatrix}$	$\begin{bmatrix} 0 \\ 0 \\ 0 \\ 1 \end{bmatrix}$
Parameter	Estimates			Null hypothesis					
Intercept	mean of 'control' group ( $\mu_1$ )			$H_0: \mu = \mu_1 = 0$					
$\alpha_2^*$	mean of group 2 minus mean of 'control' group ( $\mu_2 - \mu_1$ )			$H_0: \alpha_2^* = \mu_2 - \mu_1 = 0$					
$\alpha_3^*$	mean of group 3 minus mean of 'control' group ( $\mu_3 - \mu_1$ )			$H_0: \alpha_3^* = \mu_3 - \mu_1 = 0$					
...									

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**Single factor ANOVA**

- **Linear model**
  - $y = \beta_0 + \beta_1 X + \varepsilon$  ← Regression
  - $y = \mu + \alpha + \varepsilon$  ← ANOVA
- **Reduced model (when  $H_0$  is true,  $\alpha=0$ )**
  - $y = \beta_0 + \varepsilon$  ← Regression
  - $y = \mu + \varepsilon$  ← ANOVA
- **$H_0$ :**
  - Population group means are all equal
    - $\mu_1 = \mu_2 = \mu_3 = \dots = \mu$
    - $\alpha_1 = \alpha_2 = \dots = 0$  (Where  $\alpha_1 = \mu_1 - \mu$ )

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**Null hypotheses**

- $y = \mu + \alpha + \varepsilon \rightarrow y = \mu + \beta_1 G_1 + \beta_2 G_2 + \dots + \varepsilon$
- **Null hypotheses ( $H_0$ )**
  - Parameter based
    - Population mean = 0 ( $\mu=0$ )
    - Population slope1 = 0 ( $\beta_1=0$ )
    - Population slope2 = 0 ( $\beta_2=0$ )
    - ...
    - Use t-tests
    - But what do these slopes mean in ANOVA??

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**Null hypotheses**

- $y = \beta_0 + \beta_1 X + \varepsilon$
- **Null hypotheses ( $H_0$ )**
  - Model based (variance based)

Compare fit  $\left\{ \begin{array}{l} y = \mu + \alpha + \varepsilon \\ y = \mu + \varepsilon \end{array} \right.$

- Generate a statistic based on the ratio of fit of the full and reduced models
  - F-ratio

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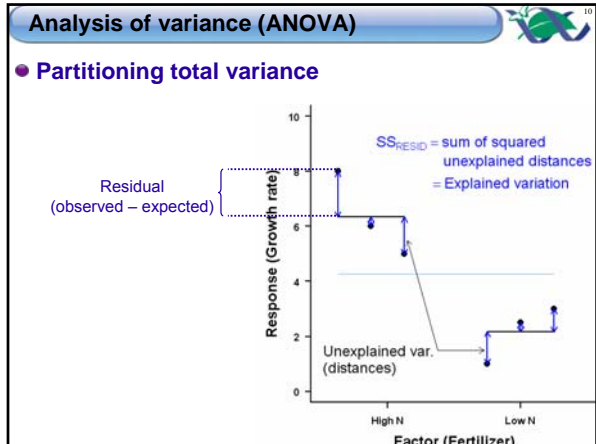
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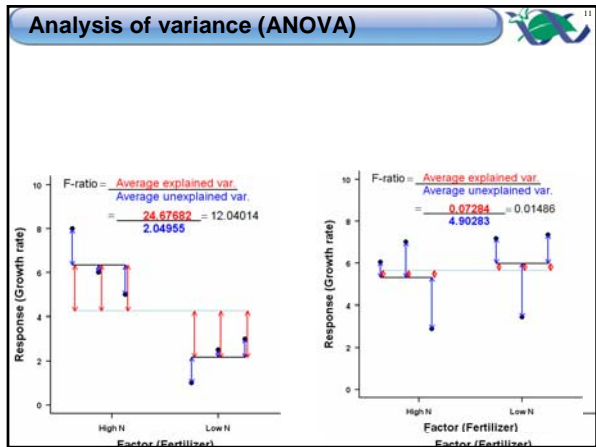
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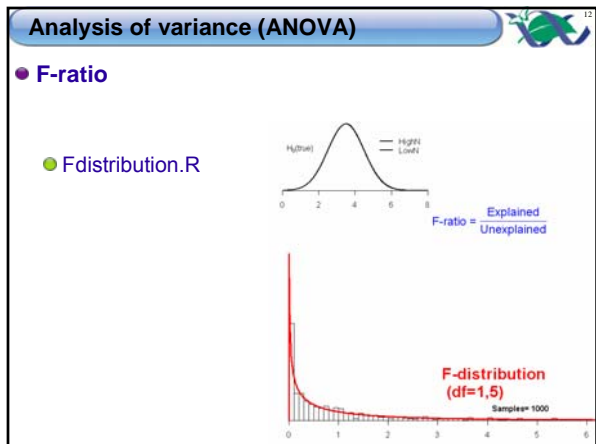
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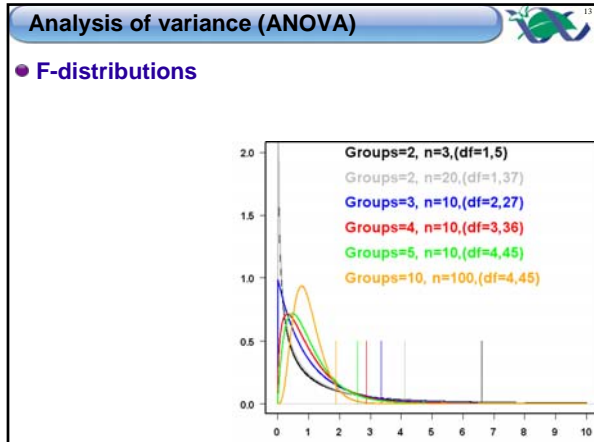
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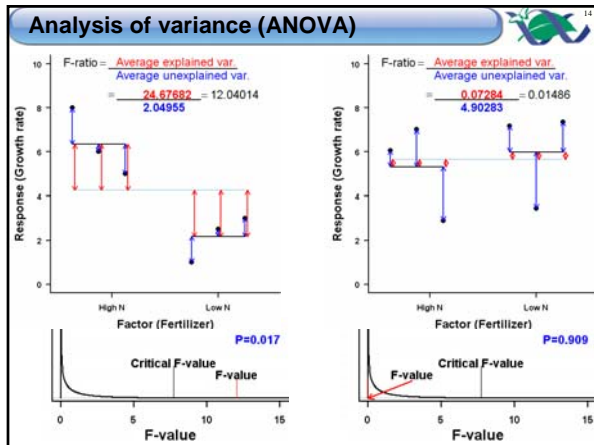
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### Analysis of variance (ANOVA)

● Weight of Lobelia seedlings grown in one of four composts (A, B, C & D)

● Biological hypothesis: seedlings grow differently in different composts

- There is an effect of compost type on the weight of Lobelia seedlings

●  $H_0: \mu_A = \mu_B = \mu_C = \mu_D = \mu$

- Population group means are all equal

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### Analysis of variance (ANOVA)

Analysis of Variance Table

Response: WEIGHT

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
COMPOST	3	29.480	9.827	6.7498	0.0009943
Residuals	36	52.411	1.456		

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- F-ratio = **Explained** / **Unexplained**
- F-ratio =  $\frac{9.827}{1.456} = 6.749$

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### Single factor ANOVA

- There is nothing on this page!
- Well why don't you delete it then?
- Why don't you get stuffed?
- Are you talking to me?
- Don't you look at me!!
- I could have a more intelligent conversation with a thimble full of pond scum!

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### Single factor ANOVA

- Assumptions
  - Independent observations
  - Normality (residuals)
    - Boxplot of response variable
  - Homogeneity of variance (residuals)
    - Spread of observations around regression line
    - Residual plot

```
> boxplot(RESPONSE ~ FACTOR, data=DATA)
```

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### Single factor ANOVA



- Fit linear model  
 $y = \mu + \alpha + \epsilon$

```
> *.lm <- lm(RESPONSE ~ FACTOR, data=DATA)
```

```
> *.aov <- aov(RESPONSE ~ FACTOR, data=DATA)
```

$$y = \mu + \alpha_2 + \alpha_3 + \alpha_4 + \epsilon$$

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### Single factor ANOVA



- Final checks (influence measures)

- Residual
  - How much each Y value differs from expected
  - Residual plot

```
> plot(*.lm)
```

OR

```
> plot(*.aov)
```

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### Single factor ANOVA



- Analysis sequence

- Design experiment/survey
- Collect data
- Test assumptions
- Fit linear model
  - Estimate parameters
  - Full vs reduced
    - Partition variability into explained & unexplained
    - $r^2$ 
      - $SS_{\text{Groups}} / SS_{\text{Total}}$

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**Single factor ANOVA**

- **Analysis sequence cont.**
  - Test  $H_0$ 's `> summary(*.aov)`
    - $\alpha_1 = \mu = 0$  (t-test)
    - $\alpha_2 - \alpha_1 = 0$  (t-test)
    - ..... (t-test)
  - Full vs Reduced (explained vs unexplained)
    - $\alpha_1 = \alpha_2 = \dots = 0$
    - F-ratio statistic =  $MS_{Groups}/MS_{Residual}$
    - F-distribution ( $df=p-1, p(n-1)$ )
- Conclusions `> anova(*.aov)`
  - Reject or not reject  $H_0$

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**Planned comparisons**

- **Following a significant ANOVA**
  - Some group comparisons more biologically meaningful than others
    - Control vs Treatment1, etc
- **Planned prior to data analysis!**
- **$p-1$  comparisons maximum**
- **All comparisons must be independent**

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**Planned comparisons**

- **Consider four groups**

$$y = \mu + \alpha_2 + \alpha_3 + \alpha_4 + \epsilon$$
  - By default, this tests
    - $\alpha_1 = \mu = 0$
    - $\alpha_2 - \alpha_1 = 0$
    - $\alpha_3 - \alpha_1 = 0$
    - $\alpha_4 - \alpha_1 = 0$
- **This might not be useful**

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**Planned comparisons**

- Consider four groups  
 $y = \mu + \alpha_2 + \alpha_3 + \alpha_4 + \varepsilon$
- We can redefine  $\alpha_2, \alpha_3, \alpha_4$ 
  - $p-1$  alternative comparisons
    - E.g.  $\alpha_2$  : group2 vs group4
    - E.g.  $\alpha_2$  : average of group1 & group2 vs group3
- Define specific comparisons
  - Need to specify which groups contribute to each effect ( $\alpha_2, \alpha_3, \alpha_4$ )

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**Planned comparisons**

- Consider four groups  
 $y = \mu + \alpha_2 + \alpha_3 + \alpha_4 + \varepsilon$
- Define specific comparisons
  - grp1 vs grp2
  - Average grp2&grp3 vs grp4

These are contrast coefficients

	Grp1	Grp2	Grp3	Grp4
$\alpha_2$	1	-1	0	0
$\alpha_3$	0	0.5	0.5	-1
$\alpha_4$	?	?	?	?

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**Planned comparisons**

- Consider four groups  
 $y = \mu + \alpha_2 + \alpha_3 + \alpha_4 + \varepsilon$
- Define specific comparisons
  - Linear trend
  - Quadratic trend
  - Cubic trend

	Grp1	Grp2	Grp3	Grp4
$\alpha_2$	-3	-1	1	3
$\alpha_3$	-1	1	1	-1
$\alpha_4$	-1	3	-3	1

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### Planned comparisons



#### Defining contrasts

```
>contrasts(FACTOR) <- cbind(c(coefs), c(coefs))
```

- Where coefs are the contrast coefficients

#### Ensure that the contrasts are independent (orthogonal)

```
>crossprod(contrasts(partridge$GROUP))
```

- All lower left (triangle) values must be 0

	[,1]	[,2]	[,3]	[,4]		[,1]	[,2]	[,3]	[,4]
[1,]	2	0	0	0	[1,]	2	0	1	0
[2,]	0	30	0	0	[2,]	0	30	-4	0
[3,]	0	0	1	0	[3,]	1	-5	4	0
[4,]	0	0	0	1	[4,]	0	0	0	1

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### Post-hoc unplanned comparisons



#### Following a significant ANOVA

- None of the possible comparisons appear more or less biologically meaningful than others
- Compare all groups pairwise
- Lots of comparisons
  - Not independent

#### Control familywise Type I error rate at 0.05

- Adjust each comparison

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### Specific comparisons of means



#### Familywise Type I error rate

- Probability of at least one Type I error among all comparisons
- Increases with increasing number of groups

No. of groups	No. of comparisons	Familywise Type I error probability
3	3	0.14
5	10	0.40
10	45	0.90

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### Post-hoc unplanned comparisons



- **Bonferroni corrections**

- Divide p-value by number of comparisons

- **Tukey's test**

- Adjust degrees of freedom according to the number of comparisons

```
>library(multcomp)
>summary(simtest(DV~FACTOR,data,type="Tukey"))
```

- **More comparisons – more brutal the adjustments**

- Reduced power

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### Factors



- **Fixed factors**

- Only interested in the specific groups (levels) tested
- Can't generalize to other levels
- E.g. control/caged
- Levels selected to represent levels interested in

- **Random factors**

- Want to extrapolate to all other possible levels
- E.g. Temperature, Site, pH
- Levels selected as random representatives

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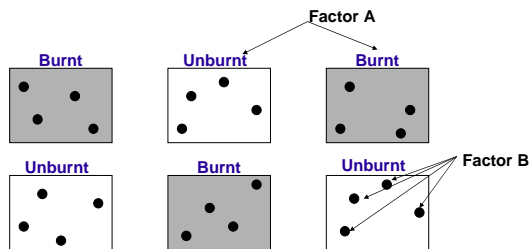
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### Nested ANOVA



- **Aims**

- Investigate the effect of a main factor (A)
- Sub-replicates collected (Factor B) within A
  - Random factor



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**Nested ANOVA**

- **Data**
  - Dependent (response variable)
  - Factor A (categorical variable)
    - Main effect
    - Fixed factor
  - Factor B (categorical variable)
    - Subreplicates
    - Random factor
    - Factor B(A)

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**Nested ANOVA**

- **Linear model**  
 $y = \mu + \alpha + \beta + \varepsilon$
- **Ho:**
  - Populations means of Factor A are equal
    - No effect of Factor A
    - $\mu_1 = \mu_2 = \dots = \mu$
    - $\alpha_1 = \alpha_2 = \dots = 0$
  - Populations means of Factor B are equal
    - No variance between levels of Factor B within Factor A
    - Not really interested in testing whether there are differences between sub-replicates!

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**Nested ANOVA**

- **ANOVA table**

	MS	F-ratio
Factor A	$MS_A$	$MS_A/MS_{B(A)}$
Factor B(A)	$MS_{B(A)}$	$MS_{B(A)}/MS_{Resid}$

The diagram illustrates a nested ANOVA design. It shows three columns representing levels of Factor A: Burnt, Unburnt, and Burnt. Each level contains three boxes representing levels of Factor B (sub-replicates). The boxes are shaded to indicate the Factor A level: Burnt (grey), Unburnt (white), and Burnt (grey). The boxes are labeled with their respective Factor A levels. Arrows point from the 'Factor A' label to the columns and from the 'Factor B' label to the boxes within a column.

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**Nested ANOVA**

Factor	MS	F-ratio
A	$MS_A$	$MS_A / MS_{B(A)}$
B(A)	$MS_{B(A)}$	$MS_{B(A)} / MS_{C(B(A))}$
C(B(A))	$MS_{C(B(A))}$	$MS_{C(B(A))} / MS_{Resid}$

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**Nested ANOVA**

- **Assumptions**
  - As with single factor ANOVA
  - Consider what the replicates are
    - Blocks are the replicates of Factor A
    - For Factor A, the average of Factor B within A are the replicates
      - These need to be used for boxplots

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**Factorial ANOVA**

- **Aims**
  - To investigate the effects of two or more factors (and their interactions) on a response variable
- **Data**
  - Dependent (response variable)
  - Factor A (categorical variable)
    - Fixed or random factor
  - Factor B (categorical variable)
    - Fixed or random factor
- **Every level of Factor A crossed with every level of Factor B**

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**Factorial ANOVA**

Factor A – White or Gray

Factor B – Stripes or not stripes

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**Factorial ANOVA**

- **Linear model**  
 $y = \mu + \alpha + \beta + \alpha\beta + \varepsilon$
- **Ho:**
  - Populations means of Factor A are equal
    - No effect of Factor A
    - $\alpha_1 = \alpha_2 = \dots = 0$
  - Populations means of Factor B are equal
    - No effect of Factor B
    - $\beta_1 = \beta_2 = \dots = 0$
  - No interaction between Factor A and Factor B

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**Factorial ANOVA**

- **ANOVA**

Factor	MS	F-ratio (both fixed)	F-ratio (A fixed, B random)	F-ratio (both random)
A	$MS_A$	$MS_A/MS_{Resid}$	$MS_A/MS_{A:B}$	$MS_A/MS_{A:B}$
B	$MS_B$	$MS_B/MS_{Resid}$	$MS_B/MS_{Resid}$	$MS_B/MS_{A:B}$
A:B	$MS_{A:B}$	$MS_{A:B}/MS_{Resid}$	$MS_{A:B}/MS_{Resid}$	$MS_{A:B}/MS_{Resid}$

- **Interpret interactions first!**
  - Simple main effects

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**Factorial ANOVA**

- **Unbalanced designs (models) - non-orthogonal**
  - Unequal sample sizes
    - Source of unequal variance
    - Complicate partitioning of variance
  - When balanced
    - $SS_{Total} = SS_A + SS_B + SS_{AB} + SS_{Residual}$
  - When unbalanced
    - $SS_{Total} \neq SS_A + SS_B + SS_{AB} + SS_{Residual}$

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**Factorial ANOVA**

- **Type I (regular, sequential) SS**
  - SS = improvement with each successive term  
 $y = \mu + \alpha + \beta + \alpha\beta + \epsilon$
  - RSS<sub>MEAN</sub>  $y = \mu + \epsilon$
  - RSS<sub>A</sub>  $y = \mu + \alpha + \epsilon$        $SS_A = RSS_{MEAN} - RSS_A$
  - RSS<sub>A,B</sub>  $y = \mu + \alpha + \beta + \epsilon$        $SS_B = RSS_A - RSS_{A,B}$
  - RSS<sub>A,B,AB</sub>  $y = \mu + \alpha + \beta + \alpha\beta + \epsilon$        $SS_{AB} = RSS_{A,B} - RSS_{A,B,AB}$
  - Order important when unbalanced  
 $y = \mu + \beta + \alpha + \alpha\beta + \epsilon$

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**Factorial ANOVA**

- **Type II (hierarchical) SS**
  - SS = improvement based on a hierarchical model comparison  
 $y = \mu + \alpha + \beta + \alpha\beta + \epsilon$
  - RSS<sub>A</sub>  $y = \mu + \alpha + \epsilon$
  - RSS<sub>B</sub>  $y = \mu + \beta + \epsilon$
  - RSS<sub>A,B</sub>  $y = \mu + \alpha + \beta + \epsilon$
  - RSS<sub>A,B,AB</sub>  $y = \mu + \alpha + \beta + \alpha\beta + \epsilon$
  - Order unimportant
  - Same as Type I !
  - $SS_A = RSS_B - RSS_{A,B}$
  - $SS_B = RSS_A - RSS_{A,B}$
  - $SS_{AB} = RSS_{A,B} - RSS_{A,B,AB}$

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### Factorial ANOVA

- **Type III (marginal, orthogonal) SS**
  - SS = improvement based comparison with and without term

$$y = \mu + \alpha + \beta + \alpha\beta + \varepsilon$$

Same as Type I !

$RSS_{B,AB}$	$y = \mu + \beta + \alpha\beta + \varepsilon$	
$RSS_{A,AB}$	$y = \mu + \alpha + \alpha\beta + \varepsilon$	
$RSS_{A,B}$	$y = \mu + \alpha + \beta + \varepsilon$	
$RSS_{A,B,AB}$	$y = \mu + \alpha + \beta + \alpha\beta + \varepsilon$	

- Order unimportant
- Best choice??

$$SS_A = RSS_{B,AB} - RSS_{A,B,AB}$$

$$SS_B = RSS_{A,AB} - RSS_{A,B,AB}$$

$$SS_{AB} = RSS_{A,B} - RSS_{A,B,AB}$$


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### Randomized block

- **Aims**
  - Groups of main treatment blocked together
  - Reduce unexplained variation without increasing sample size

Blocks (Factor B)      Factor A

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### Randomized block

- **Data**
  - Dependent variable
  - Factor A – main effect
    - Each level is observed within Factor B
  - Factor B (blocking factor)
    - Random factor

Factor A      Factor B

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**Randomized block**



- When the levels of Factor A cannot occur simultaneously – e.g. given at different times

- Called repeated measures
- Eg. Measurements taken every five minutes from the same individual

0, 5, 10, 15, 20, 30, 40	0, 5, 10, 15, 20, 30, 40
0, 5, 10, 15, 20, 30, 40	0, 5, 10, 15, 20, 30, 40

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**Randomized block**



- Linear model

$$y = \mu + \beta + \alpha + \varepsilon$$

- Ho:

- Populations means of Factor B (blocks) are equal
  - No difference between blocks
  - $\beta_1 = \beta_2 = \dots = 0$
- Populations means of Factor A are equal
  - No effect of Factor A
  - $\alpha_1 = \alpha_2 = \dots = 0$

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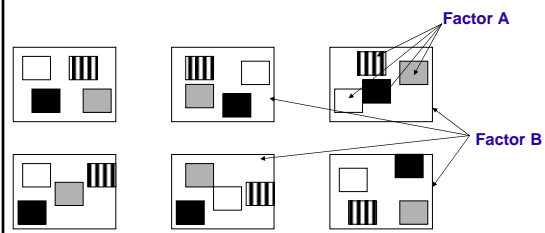
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**Randomized block**



- ANOVA table

	MS	F-ratio
Factor B	$MS_B$	$MS_B/MS_{Resid}$
Factor A	$MS_A$	$MS_A/MS_{Resid}$




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**Randomized block**

- **Assumptions**
  - Normality, equal var
  - No block (B) by A interaction
    - Reduces power of test of A
    - Difficult to interpret
    - Diagnose with an interaction plot
  - Sphericity
    - Equal variances of differences between all pairs of treatments
      - var of group1 – group2 = var of group2 – group 3 etc
    - Non sphericity chances of type I errors

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**Randomized block**

- **Sphericity likely to be met when treatments randomly assigned to units within blocks**
  - E.g. treatments randomized within quadrats
  - E.g. order of drugs given to rats random
- **Sphericity not likely to be met when treatments are time**
  - Since cant randomize time
  - Time 1 is likely to be more similar to Time 2 than to Time 10
- Calculate the degree to which sphericity not met
  - Epsilon (<1)
    - Greenhouse Geisser (preferred)
    - Huyhn-Feldt
  - Adjust degrees of freedom

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**Split-plot designs**

- **Aims**
  - Combination of nested and blocking designs

The diagram illustrates split-plot designs with three factors: Factor A, Factor B, and Factor C. It shows a 2x2 grid of plots. The top-left plot is a 2x2 grid of treatments (white, black, grey, striped) with Factor A indicated by a large arrow. The top-right plot is a 2x2 grid of treatments with Factor C indicated by a large arrow. The bottom-left plot is a 2x2 grid of treatments with Factor B indicated by a large arrow. The bottom-right plot is a 2x2 grid of treatments with Factor B indicated by a large arrow. Arrows from the labels 'Factor A', 'Factor B', and 'Factor C' point to their respective plots in the grid.

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**Split-plot designs**

- **Data**
  - Factor A (between plot effect)
  - Factor B (plot)
  - Factor C (within plot effect)

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**Split-plot designs**

- **Data**
  - Factor A (between plot effect)
    - Main effect 1
  - Factor B (plot)
    - Random factor
  - Factor C (within plot effect)
    - Main effect 2

} **Nested**

} **Blocking**

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**Split-plot designs**

- **ANOVA**

Factor	MS	F-ratio
A	$MS_A$	$MS_A/MS_B$
B(A)	$MS_{B(A)}$	$MS_{B(A)}/MS_{Resid}$
C	$MS_C$	$MS_C/MS_{B(A):C}$
A:C	$MS_{A:C}$	$MS_{A:C}/MS_{B(A):C}$
B(A):C	$MS_{B(A):C}$	$MS_{B(A):C}/MS_{Resid}$

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## Split-plot designs



### ● Assumptions

- Normality, equal variance
  - Establish correct residuals
- Sphericity (for within plot effects)
  - As for blocking designs
- No interaction between plot and within plot effects

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